# POPULATION CONNECTIVITY IN LAKE MICHIGAN YELLOW PERCH 

by

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Dedicated to my parents.
Thank you for supporting every foolish venture.

And to Gary, obviously.

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#### Abstract

Yellow perch (Perca flavescens) are an ecologically and economically important species in the Great Lakes. In Lake Michigan, there is substantial interest in restoring Yellow perch to their historic abundance. Populations in Lake Michigan began to decline dramatically in the late 1980s and, despite management efforts, have not rebounded. Understanding stock structure is imperative for implementing successful management strategies, but assessing population structure in yellow perch is particularly challenging. Yellow perch disperse on surface currents during a 30-40 day larval period where dispersal is difficult to observe directly and varies interannually. In order to better understand yellow perch population structure and connectivity, we sequenced 960 individuals, collected at 20 sampling locations across Lake Michigan using restriction siteassociated DNA sequencing (RADseq). We used 3337 single nucleotide polymorphisms (SNPs) to observe genetic differences between populations and paired these findings with a Lagrangian particle tracking model to explain patterns of connectivity and gene flow. We showed that Green Bay and the main basin of Lake Michigan are genetically distinct populations. Within these two genetic groups, drivers for population structure appear to be very different. Green Bay shows distinct populations across its 150 kilometers, consistent with isolation by distance. These populations show lower global allelic richness and heterozygosity than the main basin. In contrast, the main basin shows low but significant genetic distance (measured as pairwise $\mathrm{FST}_{\text {S }}$ ) and higher allelic richness and observed heterozygosity, appearing to behave more like a marine system. To validate this observation, we paired these findings with a Lagrangian particle tracking model to explain patterns of connectivity and gene flow and found that distances derived from these particle tracking models were significantly correlated with the genetic distances observed between main basin populations.


## CHAPTER 1. POPULATION CONNECTIVITY IN LAKE MICHIGAN YELLOW PERCH

### 1.1 Introduction

Understanding patterns of population connectivity is imperative for implementing successful species conservation and management strategies (Allendorf et al., 2010; Fraser \& Bernatchez, 2001). In many fishes, dispersal plays a fundamental role in structuring populations, but quantifying the direction and magnitude of dispersal can be challenging (Hedgecock et al., 2007). Many freshwater fishes and over $95 \%$ of all marine fishes have a pelagic larval stage (Nelson 2006). Larval fish can disperse vast distances-sometimes hundreds of kilometers from where they've spawned-on surface currents (Cowen \& Sponaugle, 2009). Tracking these larvae, which are often small and transparent, is not feasible using traditional sampling methods (Bradbury \& Snelgrove, 2001; Levin, 2006) and, because larval transport cannot be observed in situ, integrative approaches are necessary to reconstruct how individuals move and interact across space and time (Burgess et al., 2014; Weersing \& Toonen, 2009) Genomic tools allow us to look at patterns of connectivity post-dispersal and identify gene flow among populations. In doing so, we can better define management units and develop informed strategies for population recovery and persistence (Bernatchez et al., 2017; Johnson et al., 2018). However, while genomic tools can reveal discrete population structure, it cannot explain what drives population structure. In order to relate pattern to process, researchers pair other analytical approaches with genetics (Selkoe et al., 2016). These multi-pronged approaches are common in marine systems and include otolith microchemistry (Collins et al., 2013; Feyrer et al., 2007; Hoey et al., 2020), investigating boundaries based on environmental features such as depth, turbidity, and temperature (Johansson et al., 2015; Lawlor \& Arellano, 2020; Roy et al., 2012), and modeling currents to estimate connectivity (Alberto et al., 2011; White et al., 2010). While the Great Lakes are a freshwater system, in many ways, it can behave similarly to a marine environment. Many fishes in the Great Lakes also undergo a larval dispersal stage. Moreover, due to their size, Great Lakes' physical processes occur on spatial and temporal scales more comparable to the coastal ocean (Beletsky et al., 1999), making it an appealing system for applying the aforementioned techniques.

Yellow perch (Perca flavescens) are an ecologically and economically important Great Lakes fish. Ecologically, they play an important role in nutrient cycling in the near-shore environment. They are an abundant near-shore species in many areas of the Great Lakes, serve as important predators of small fishes and invertebrates, and are an important prey item for species such as walleye (Sander vitreus), northern pike (Esox Lucius) and lake trout (Salvelinus namaycush) (Evans, 1986). Historically, yellow perch supported a lucrative commercial fishery in Lake Michigan with catches reaching 1.1 million kg annually. Yellow perch have also been a popular fish for recreational anglers-at the peak of this fishery, yellow perch comprised $85 \%$ of recreational harvest (Francis et al., 1996). Yellow perch populations began to decline in the late 1980's and early 1990's, which prompted the complete closure of commercial fisheries save for Green Bay and Grand Traverse Bay, reduced bag limits and seasonal spawning closures (Clapp \& Dettmers, 2004). Despite these management efforts yellow perch populations have not rebounded. Annual surveys conducted across management jurisdictions and reported by the Lake Michigan yellow perch task force show that the current adult abundance remains well below historically observed abundance in all areas of the lake (GLFC Lake Michigan Committee, 2018). The cause of yellow perch decline in Lake Michigan is not fully understood and likely has many contributing factors including invasive dreissenid mussels filtering out nearshore nutrients (Goto et al., 2020; Hecky et al., 2004; Mida et al., 2010), competition with and predation by invasive alewife (Alosa pseudoharengus) (Evans, 1986; Janssen \& Luebke, 2004), and low adult abundance (i.e., spawning stock biomass) following overexploitation (Wilberg et al., 2005). Annual surveys of yellow perch show high variability in recruitment and year class strength, and low production across year and region (GLFC Lake Michigan Committee, 2018). Studies suggest that recruitment might be limited by declining zooplankton abundance (Dettmers et al., 2003), larval predation (Madenjian et al., 2008), and adult abundance (Clapp \& Dettmers, 2004; Wilberg et al., 2005). However, nearshore dynamics might only partially explain interannual variability in yellow perch recruitment.

Similar to many marine fishes, yellow perch have a pelagic larval stage that lasts between thirty and forty days (Whiteside et al., 1985). Especially during their early development, yellow perch are poor swimmers (Dettmers et al., 2005) and are likely to be transported by surface currents. The dynamic circulation patterns found in Lake Michigan coupled with the yellow perch pelagic larval duration, likely influences recruitment success. Previous studies have used biophysical models to predict the extent of larval dispersal throughout Lake Michigan. Beletksy et
al., (2007) suggested that summer current patterns could result in larval yellow perch dispersing from western to eastern shores but are unlikely to result in movement between the southern and northern basins. Because Lake Michigan sits on a steep latitudinal gradient (494 km north to south), lack of mixing could lead to genetically distinct, locally adapted populations. Additionally, Lake Michigan's two large bays, Green Bay and Grand Traverse Bay, (Figure 1.1a), often experience their own, weaker, regional currents and mix very little with the main basin (Beletsky \& Schwab, 2008) and this could present a significant barrier to gene flow.

Previous genetic studies have suggested that the yellow perch in the Great Lakes share an evolutionary history stemming from a single colonization event, with most populations across the Great Lakes sharing a single mtDNA haplotype (Sepulveda-Villet et al., 2009). However, additional studies have shown complex population structure in yellow perch in Lake Erie (Sepulveda-Villet \& Stepien, 2011) and there is no reason to anticipate population dynamics in Lake Michigan would be any less complex. A microsatellite study on yellow perch in Lake Michigan (Miller, 2003) revealed moderate genetic structure between populations in Green Bay and the southern region of the main basin. These populations were found to be genetically similar, which matches predictions made with Beletsky et al.'s (2007) biophysical model. However, this study did not include populations from the northern portion of Lake Michigan. Without sampling geographically representative populations throughout the entire basin, it is challenging to fully understand how dispersal might affect gene flow across the entirety of the lake.

Here, we build on previous work by sequencing 960 yellow perch from 20 geographically representative locations across Lake Michigan (Figure 1.1a). Using restriction site associated DNA sequencing (RADseq) we identified population structure across the lake and paired these results with Lagrangian particle tracking model. By coupling genetic data with biophysical processes we were able to better understand how lake currents influence larval dispersal patterns, and more broadly, population connectivity across the lake.

### 1.2 Materials and Methods

### 1.2.1 Sample Collection

Yellow perch were collected from 20 sites circumscribing Lake Michigan and Green Bay in 2018 and 2019 (Figure 1.1a, Table 1).

In 2018, adult individuals, defined as year one or older ( $>100 \mathrm{~mm}$ ), were sampled using 12 hour overnight multi-mesh gill net sets. Net location and depth varied by site and were informed by agency correspondence and local fishing reports. Sampling during the 2018 season began in southern Lake Michigan in early March and continued through the end of July at the northernmost sites in order to correspond with regional spawning times. Young of year ( $<100 \mathrm{~mm}$ ) were sampled in September using beach seines. The 2019 sampling regime was greatly assistance by several agencies-the Michigan Department of Natural Resources, the Indiana Department of Natural Resources, the Wisconsin Department of Natural Resources, the Illinois Natural History Survey, and the Grand Traverse Band of Ottowa and Chippewa. Agency sampling methods consisted of bottom trawls, gill netting, and creel surveys, and took place across both 2018 and 2019. 1376 individuals were sampled over two sampling seasons, and a subset of 960 were selected for sequencing (Table 1.1). Tissues were stored at $-4^{\circ} \mathrm{C}$ upon arrival at Purdue University and held until DNA extraction.

### 1.2.2 DNA extraction

DNA was isolated from fin tissue with Qiagen DNeasy® Blood \& Tissue Kits. Following an overnight ( $\sim 14$ hour) tissue digestion with Proteinase K incubated at $56^{\circ} \mathrm{C}$, DNA was extracted in plates using standard kit protocols, and eluted from the silica membrane with 200uL Tris LowEDTA buffer. Extractions were quantified using a Quant-it ${ }^{\text {TM }}$ PicoGreen® dsDNA Assay (Invitrogen, Waltham, MA), and DNA was normalized to a quantity of 200ng or approximately $20 \mathrm{ng} / \mu \mathrm{L}$.

### 1.2.3 RAD library prep and sequencing

Libraries for restriction site-associated DNA (RAD) sequencing were prepared following the BestRAD protocol (Ali et al., 2016). Normalized DNA was digested with the restriction enzyme

SbfI followed by ligation with indexed adaptors. Barcoded libraries were pooled into master libraries of 96 individuals and fragmented to $\sim 300-500 \mathrm{bp}$ with 12 30s cycles in a Q500 sonicator (Qsonica, Newtown, CT). Fragmented DNA was bound to Dynabeads ${ }^{\text {TM }}$ M-280 Streptavidin magnetic beads (Invitrogen) and washed with buffer to remove non-target fragments. Following purification with AMPure XP beads (Beckman Coulter, Brea, CA), master libraries were input into the NEBNext ${ }^{\circledR}$ Ultra ${ }^{\text {TM }}$ DNA Library Prep Kit for Illumina ${ }^{\circledR}$ at the End Prep step for ligation of master library barcodes, a 250-bp insert size-selection, and a 12-cycle PCR enrichment. Successful size-selection and enrichment were confirmed with visualization of products on a $2 \%$ agarose E-Gel (Invitrogen). Products underwent a final AMPure XP purification clean-up followed by quantification with a Qubit® 2.0 Fluorometer. A total of 10 master libraries, each containing 96 individually barcoded samples, were sent to Novogene (Sacramento, CA) for PE150 sequencing on one lane of the Illumina NovaseqS4 platform.

### 1.2.4 Read processing and SNP filtering

Raw Illumina RAD sequence reads were processed using the STACKS v2.54 (Rochette et al., 2019) software pipeline. Reads were cleaned and demultiplexed by barcode using the STACKS subprogram process_radtags. Sequences were demultiplexed by barcode, filtered for illumina quality score and enzyme cut-site, and trimmed to 140 base pairs to reduce tail-end sequencing errors (parameter flags $=-$-filter_illumina, --bestrad, -t 140). The resulting filtered, individually assigned reads were aligned to the yellow perch reference genome ( $P$. flavescens PFLA_1.0 assembly, GenBank accession GCA_004354835.1; Feron et al., 2020) with bowtie2 (Langmead et al., 2019, Langmead \& Salzberg, 2012) (parameter flag = --very-sensitive). SNPS were called from reference aligned paired end reads with the STACKS subprogram gstacks (parameter flag = --rm-unpaired reads) and individuals were genotyped at each identified SNP. The gstacks output files, which contain consensus sequences at each loci identified as well as individual genotyping data, were filtered through STACKS subprogram populations. SNPS that genotyped in less than $30 \%$ of individuals were discarded (parameter flags $=-r 0.3$ ) and results were exported in variant call format (VCF). In a final filtering step, SNPs were filtered using vcftools v0.1.9 (Danecek et al., 2011). Filtering consisted of first removing SNPs that genotyped in fewer than $80 \%$ of individuals, then removing outlier SNPs with regard to mean read depth across individuals (--minmeanDP 10, --max-meanDP 60) as recommended in O’Leary et al., (2018). In the final filtering
step, individuals with $>50 \%$ missing data were identified and removed (mean $\%$ missingness was 7.7\%). The resulting vcf file was converted to GENEPOP and STRUCTURE format for downstream analysis using PGDSpider (Lischer \& Excoffier, 2012).

### 1.2.5 Summary statistics

Population level summary statistics including observed heterozygosity (Ho), expected heterozygosity $(H e)$, and allelic richness $(A r)$ were calculated using the in the R package hierfstat v0.04.22 (Goudet, 2005), which returns loci and population level statistics using equations as defined in Nei, 1987. A Bartlett test of homogeneity of variances and a paired $t$-test between Ho and He was performed on each population to determine whether or not observed heterozygosity significantly differed from expected heterozygosity. ANOVAs were used to determine whether global genetic diversity estimates (Ho and $A r$ ) between Green Bay and the Main basin statistically differed from one another.

### 1.2.6 Identifying population structure

A Bayesian clustering method was implemented in STRUCTURE v2.3.4 (Pritchard et al., 2000) to determine population structure present among all populations. STRUCTURE was also used to examine potential fine-scale structure in the main basin, as the populations in the main basin did not follow any discernible pattern of isolation by distance, nor show any demonstrable structure in either principal components or principal coordinates analysis. For STRUCTUE analysis with all 26 populations, optimal K was determined using the delta K method (Evanno et al., 2005). Runs consisted of an initial burn in period of 50,000 Markov Chain Monte Carlo (MCMC) iterations followed by 50,000 iterations for each inferred cluster. Analyses were performed with $\mathrm{K}=1-30$ clusters and replicated five times for each K. For main basin STURCTURE analyses, we employed admixture and correlated allele frequency models, as they are considered to be most appropriate when subtle population structure is expected (Falush et al., 2003; Hubisz et al., 2009). Analyses were performed for $K=1-25$ and repeated five times for each $K$.

Pairwise $\mathrm{F}_{\text {ST }}$ between populations and $95 \%$ confidence intervals were calculated in hierfstat v0.5.7 (Goudet, 2005). Non-zero pairwise FST's based on confidence intervals were reported in Table 1.2 and exported for analysis in GenAlEx v6.5 (Peakall \& Smouse, 2006, 2012)
where principal coordinate analysis was performed for the in order to visualize structure among populations (Figure 1.1b).

### 1.2.7 Isolation by distance

Pairwise $\mathrm{F}_{\text {ST }}$ and Euclidean distances (defined as the straight-line distance between two geographic points) was compared using non parametric permutational testing (Smouse \& Peakall, 1999). Because both genetic and geographic distance matrices are not independent, the assumptions of traditional linear regressions can be violated. Instead, we shuffled the y matrix (genetic distance) and assessed whether the relationship between the true matrices (Rxy) varies significantly from the relationship between random shuffles. A total of 999 randomization permutations were performed in order to estimate what values would be anticipated if the matrices were unrelated and these values were used to calculate the modified P value (Peakall \& Smouse, 2012).

### 1.2.8 Oceanographic Distance

For the biophysical model, we used a Lagrangian particle tracking model developed previously to study transport of larval cod (Churchill_et_al.2__2011;_Huret_et_al.,__2007, http://fvcom.smast.umassd.edu/). Vertical mixing was implemented using turbulent diffusivity output from FVCOM and a random-walk scheme for spatially varying vertical diffusivity. Rowe et al. (2016) recently modified the code to include a vertical floating/sinking/swimming velocity (Gräwe, 2011). Particles were assumed to be neutrally buoyant in the preliminary simulation, but a vertical swimming velocity of 0.0003 was also included. The Lagrangian particle tracking simulations were forced by output from a FVCOM simulation of Lake Michigan-Huron developed by Anderson and Schwab (E. J. Anderson \& Schwab, 2013) to simulate currents in the Straits of Mackinac. Horizontal grid resolution varied with finer resolution nearshore and in regions with complex coastlines ( 100 m in the Straits to 2.5 km in the center of the lakes), and each horizontal grid included 20 sigma layers in the vertical. The FVCOM is an unstructured grid, finite-volume, free surface, three-dimensional primitive equation ocean model that solves the momentum, continuity, temperature, salinity, and density equations (Chen et al., 2003). The unstructured grid of FVCOM conforms to complex coastline morphologies and allows for increased grid resolution in regions of interest. Turbulence closure was implemented through the MY- 2.5 scheme for
vertical mixing (Galperin et al., 1988), and the Smagorinsky scheme for horizontal mixing (Smagorinsky, 1963).FVCOM has been implemented for the Great Lakes yielding accurate predictions of temperature, water levels, and currents (E. Anderson et al., 2010; E. J. Anderson et al., 2015; E. J. Anderson \& Schwab, 2013; Bai et al., 2013) and is being used for NOAA's nextgeneration Great Lakes Operational Forecast System (GLOFS).

To generate a connectivity matrix, the probability of transport from region $i$ to region $f$ was calculated as Nif/Ni, where Nif is the number of particles initiated in region ithat are within region $f$ at the end of the simulation, and $N i$ is the total number of particles that were initiated in region $i$. Connectivity matrices were developed for six years (2014-2019). The sensitivity of the connectivity matrices to model assumptions can be evaluated by considering scenarios of, 1.) swimming behavior, and 2.) horizontal diffusion (Table 1.3). Simple behavior scenarios were tested, including passive particle, upward swimming, and downward swimming. Realistic estimates of swimming velocity were implemented as a deterministic vertical velocity in the vertical random walk turbulence scheme; thus representing the combined effects of turbulence and directed swimming. Scenarios including horizontal diffusion were tested, using literature values (Okubo, 1971), and using existing data from drifter deployments in Lake Michigan-Huron (www.nefsc.noaa.gov/drifter) resulting in a final horizontal diffusion coefficient of $5.6 \mathrm{~m}^{2} / \mathrm{s}$. These models were run for a larval duration of 50 days, with bi-weekly release dates ranging from late May to late July (Table 1.3).

Because some sampling sites fell within the same region utilized for estimating particle transport, the dynamics of transport between them could not be resolved. Thus, we selected one population from each region for analyses pairing oceanic and genetic distance.

Lagrangian particle transport models are inherently asymmetrical, allowing reciprocal exchange between populations where the probability of transport from point $a$ to point $b$ might have a different value than the probability of transport from point $b$ to point $a$. To match our pairwise $\mathrm{F}_{\text {ST }}$ matrices, triangular distance matrices were created from asymmetrical, square matrix model outputs by averaging across the upper and lower probability matrices.

Models were generated for multiple release dates (May through July) from 2014 to 2019. Matrices based on mean triangular distance matrices were generated for each release date across
the six years for which the models were run. We tested the explanatory power of these models by regressing them with the genetic distance (pairwise $\mathrm{F}_{\mathrm{ST}}$ ) using the same statistical methods as the Euclidean isolation by distance analysis.

### 1.3 Results

### 1.3.1 Sequencing and genotyping

960 individuals were RAD sequenced, producing over 5 billion reads that resulted in an average of 5,463,100 paired-end reads per sample. Following filtering, 935 individuals represent from 26 sampling sites, year, and age class were genotyped at 3337 loci (Table .1). Mean read depth of loci across individuals was 29 x and mean missingness per individual was $7.7 \%$.

### 1.3.2 Summary statistics

The estimates of global genetic diversity for heterozygosity and allelic richness were $H o=0.449$, $H e=0.451$, and $A r=1.958$, respectively. Genetic diversity estimates were also calculated for each population (Table 1). In order to better understand processes driving genetic differentiation between Green Bay and the main basin, an ANOVA was run to determine if Ho and Ar varied significantly between the main basin and Green Bay. Ho was not found to be significantly different between groups, $(p=0.06)$. However, allelic richness was significantly different between Green Bay and the main basin ( $\mathrm{p}<0.0001$ ), with values ranging from 1.490-1.939 for Green Bay and $1.959-1.973$ for the main basin (Table 1).

### 1.3.3 Identifying population structure

Both mean likelihood values $(\mathrm{L}(\mathrm{K}))$ and $\Delta \mathrm{K}$ suggested two clusters $(\mathrm{k}=2)$ across the entire basin. This clustering reveals a distinct genetic split between Green Bay populations and Main basin samples (Figure 1.1c). When the main basin samples were run separately to determine fine-scale population structure, analysis revealed no optimal K, suggesting minimal population structure in across Main basin samples.

Populations in Green Bay, particularly south and central, are largely isolated from the main basin. The only connection to the main basin is at the northern end of that bay, about 100 km from
the southernmost site sampled. There is little exchange of current into and out of the bay (Beletsky \& Schwab, 2008), suggesting that larval recruits are unlikely to be exchanged between Green Bay and the main basin. The STRUCTURE results (Figure 1.1c) show a larger proportion of Green Bay ancestry (blue) in the southern and central populations, supporting the idea of decreasing exchange rates of individuals with increasing distance from the mouth of the bay. To quantify this, we regressed proportion Green Bay ancestry with distance from the mouth of Green Bay and a found significant linear relationship between proportion main basin ancestry and distance from the mouth (Figure 1.2b). Distance to mouth was calculated by using the straight-line distance between population coordinates and a point in the center of the mouth of Green Bay.

### 1.3.4 Isolation by distance

There is a strong linear relationship between Fst and Euclidean distances in the Green Bay dataset $\left(R^{2}=0.6213\right.$, Mantel $\left.r=0.788, p=0.004\right)$. Inversely, the Main basin dataset showed no relationship between Fst and Euclidean distance ( $\mathrm{R}^{2}=0.0009$, Mantel $\mathrm{r}=0.030, \mathrm{p}=0.363$ ). However, a reduced, linearized $\mathrm{F}_{\text {ST }}$ matrix did correlate with the probability matrix derived from oceanographic current $(\mathrm{Rxy}=-0.499, \mathrm{p}=0.014)$. The model averaged from 2014-2019 for the mid-June release dates had a significant relationship to the genetic connectivity matrix for the main basin, showing the potential for oceanographic data to resolve fine scale patterns of population connectivity (Figure 1.3b).

### 1.4 Discussion

Comparisons of population structure show a distinct genetic split between Green Bay and the Main basin, supporting previous genetic analyses of yellow perch in Lake Michigan (Miller, 2003). When the Main basin and Green Bay are inspected separately, two different patterns of genetic differentiation emerge. Green Bay shows clear population differentiation between sites that fit a traditional model of isolation by distance. Sites across Green Bay cluster spatially (Fig 1B), and vary little cross years (SGB18, SGB19) and age classes (BNDYO, BND19 and LBDYO, LBD19). This spatial clustering is supported by the strong linear relationship between genetic distance and Euclidean distance (Figure 1.2a). Alternatively, the Main basin population structure appears to be panmictic with genetic differentiation remaining relatively low between sites and no significant
relationship to Euclidean distance (Figure 1.3a). Overall, Green Bay showed a much stronger signal for population differentiation (global Fst $=0.010$ ) compared to the Main basin (global Fst $=0.0031$ ), over a much smaller geographic range.

Summer currents in Green Bay are relatively weak, particularly compared to Main basin currents (Beletsky \& Schwab, 2008). A weak, anticyclonic gyre often forms in northern Green Bay, allowing for some mixing with the Main basin, while the south remains relatively isolated (Beletsky et al., 1999; Beletsky \& Schwab, 2001). Adult yellow perch have modest home ranges and high spawning site fidelity (Glover et al., 2008), suggesting that they do not disperse after larval settlement. The patterns observed in Green Bay, i.e. strong genetic partitioning based on geographic distance in the absence of strong current, support the idea that larval dispersal is the primary driver of population connectivity in Lake Michigan. The increased admixture at northern Green Bay sites (quantified by proportion ancestry) track with the current models, which show slight mixing in the northern portion of Green Bay that does not reach the central and southern sites.

While Miller (2003) did not detect genetic differentiation between southern main basin populations, we were able to detect low but significant population structure in the main basin. This, coupled with high allelic richness (global Fst $=0.0031, \mathrm{p}=0.004$; mean $\mathrm{Ar}=1.965$ ) is similar to patterns of genetic diversity in marine fishes (DeWoody \& Avise, 2000; Martinez et al., 2018). While STRUCTURE was unable to detect population structure among Main basin populations, and traditional isolation by distance yielded no significant correlation, patterns of connectivity were correlated with derived oceanic distances. In this scenario, the ability of oceanographic distances to resolve genetic isolation by distance patterns demonstrates the value of considering dispersal in estimating population connectivity. Additionally, the juxtaposition between population structure in Green Bay and the Main basin underscore the necessity of considering current as a driver for population structure in species with pelagic larval dispersal stages in large lakes.

While our derived ocean distance explained some of the genetic differentiation amongst main basin populations, it is likely that myriad additional environmental factors play into population dynamics of yellow perch across Lake Michigan. At higher dispersal levels, it becomes
increasingly difficult to detect barriers of dispersal, arising from the system becoming more panmictic and populations less differentiated (Wang, 2013). Our models, which broadly estimate dispersal patterns across multiple years, likely do not capture fine scale environmental heterogeneity that could influence annual variance in recruitment, survival, and admixtureparticularly on an annual basis (Nanninga et al., 2014; Riginos \& Liggins, 2013). Moreover, along Lake Michigan's steep longitudinal gradient, the environment varies quite dramatically. Water temperature, substrate composition, anthropogenic disturbance, and many other environmental factors may influence local adaptation. Studies have shown that, despite near panmixia, many marine species show signatures of local adaptation across the species range (Clarke et al., 2010; Hoey \& Pinsky, 2018; Wilder et al., 2020). While SNPs used in this analysis were putatively neutral, identifying adaptive loci in this dataset could uncover further population differentiation.

Our study provides the first comprehensive, lake-wide investigation of yellow perch in Lake Michigan. We identified two regionally differentiated populations, with unique drivers for population structure: isolation by distance, and isolation by environment. This study provides the foundation for better understanding the ecology and evolution of yellow perch in Lake Michigan, which is vital for meeting conservation management goals, and rebuilding declining stocks.

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### 1.6 Figures and Tables

Figure 1.1. (a) A map of sample locations across Lake Michigan including Green Bay (orange) and the main basin (blue). (b) Principal Coordinates Analysis showing relatedness between



individual populations based on pairwise fst. (c) Results of individuals clustering analysis with
STRUCTURE ( $\mathrm{K}=2$ ). Colored lines correspond to an individual's estimated proportion membership to each cluster, and individuals are grouped by population (bold black lines).

Table 1.1. Summary statistics for all populations ( $\mathrm{n}=26$ ). Number of individuals sampled ( N ) and summary statistics including observed heterozygosity (Ho), expected heterozygosity ( He ), allelic richness, ( Ar ), and $\%$ missing data.

|  | Site ID | N | Ho | He | $A r$ | \%missing |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Main Basin |  |  |  |  |  |  |
|  | MIL19 | 32 | 0.491 | 0.451 | 1.963 | 7.78 |
|  | WAK18 | 20 | 0.490 | 0.446 | 1.964 | 6.32 |
|  | NCH18 | 15 | 0.464 | 0.439 | 1.960 | 10.82 |
|  | CHI18 | 12 | 0.481 | 0.442 | 1.967 | 2.68 |
|  | SCH18 | 14 | 0.480 | 0.442 | 1.963 | 8.44 |
|  | MIC18 | 49 | 0.494 | 0.454 | 1.963 | 5.77 |
|  | MIC19 | 46 | 0.498 | 0.452 | 1.962 | 8.57 |
|  | MICYO | 28 | 0.497 | 0.452 | 1.964 | 4.26 |
|  | STJ18 | 49 | 0.504 | 0.452 | 1.961 | 10.57 |
|  | GRH18 | 50 | 0.487 | 0.454 | 1.962 | 4.04 |
|  | GRH19 | 38 | 0.638 | 0.459 | 1.968 | 10.32 |
|  | SOH18 | 17 | 0.502 | 0.445 | 1.963 | 5.12 |
|  | LUD18 | 8 | 0.484 | 0.425 | 1.959 | 7.14 |
|  | SUT18 | 32 | 0.498 | 0.457 | 1.970 | 7.82 |
|  | NPT18 | 34 | 0.518 | 0.462 | 1.973 | 7.56 |
|  | CHX19 | 59 | 0.512 | 0.460 | 1.969 | 10.01 |
|  | CHE18 | 40 | 0.494 | 0.459 | 1.969 | 3.91 |
|  | NUB18 | 58 | 0.494 | 0.459 | 1.968 | 9.92 |
|  | MAN18 | 14 | 0.478 | 0.449 | 1.970 | 5.83 |

Table 1.1 continued

| Green Bay |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | BDN19 | 39 | 0.480 | 0.439 | 1.949 | 8.41 |
|  | BDNYO | 41 | 0.476 | 0.437 | 1.948 | 7.55 |
|  | SGB18 | 60 | 0.451 | 0.421 | 1.927 | 5.46 |


| SGB19 | 50 | 0.509 | 0.426 | 1.931 | 12.50 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| LBD19 | 40 | 0.464 | 0.433 | 1.941 | 7.59 |
| LBDYO | 41 | 0.469 | 0.432 | 1.939 | 7.56 |
| MEN19 | 49 | 0.456 | 0.425 | 1.932 | 7.00 |

Table 1.2. Pairwise Fst's across all populations. Values that did not differ from zero based on $95 \%$ CI's were excluded.

| BDN19 | BDNYO | CH118 | WAK18 | NCH18 | SCH18 S | SGB18 | SGB19 | GRH18 | GRH19 | MIC18 | MIC19 | LBD19 | LBDYO | MAN18 C | CHE18 | CHX19 | miço | MEN19 | MIL19 | STJ18 | LUD18 | SOH18 | SUT18 | NPT18 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.000886 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.080302 | 0.083696 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.085694 | 0.087729 | 0.001972 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.085058 | 0.088456 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.087246 | 0.090529 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.01297 | 0.014286 | 0.107875 | 0.114378 | 0.113863 | 0.118692 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.013243 | 0.013619 | 0.106315 | 0.109853 | 0.110386 | 0.116292 | 0.002657 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.089172 | 0.091631 | 0.001657 | 0.001534 | 0.001486 |  | 0.11734 | 0.113624 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.082003 | 0.084792 | 0.008964 | 0.006318 | 0.008144 | 0.006484 | 0.110203 | 0.106388 | 0.002534 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.085971 | 0.088534 | 0.002526 | 0.001225 | 0.001511 |  | 0.113945 | 0.11015 |  | 0.002616 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.088336 | 0.091027 |  | 0.001155 |  |  | 0.117297 | 0.113105 |  | 0.002238 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.007335 | 0.00768 | 0.090135 | 0.094699 | 0.095151 | 0.097635 | 0.017116 | 0.01273 | 0.099016 | 0.092742 | 0.095549 | 0.097809 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.010976 | 0.010406 | 0.094632 | 0.097397 | 0.098045 | 0.099958 | 0.020031 | 0.013335 | 0.101969 | 0.095274 | 0.098526 | 0.101756 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.064429 | 0.06662 |  | 0.004452 | 0.003786 | 0.002169 | 0.093669 | 0.090398 | 0.003858 | 0.008334 | 0.003755 | 0.003675 | 0.074065 | 0.07783 |  |  |  |  |  |  |  |  |  |  |  |
| 0.072846 | 0.075471 | 0.002853 | 0.003416 | 0.002814 | 0.001617 | 0.101838 | 0.098084 | 0.002457 | 0.003947 | 0.002692 | 0.002821 | 0.08284 | 0.085788 |  |  |  |  |  |  |  |  |  |  |  |
| 0.069283 | 0.071188 | 0.004913 | 0.004913 | 0.00501 | 0.003568 | 0.097372 | 0.092875 | 0.004236 | 0.003619 | 0.00337 | 0.004543 | 0.0787 | 0.081841 |  | 0.001165 |  |  |  |  |  |  |  |  |  |
| 0.08506 | 0.087524 | 0.001621 |  | 0.003182 |  | 0.115313 | 0.111367 |  | 0.003682 | 0.000796 | 0.001112 | 0.095975 | 0.099727 |  | 0.001625 | 0.0033 |  |  |  |  |  |  |  |  |
| 0.012263 | 0.012718 | 0.104393 | 0.108594 | 0.10871 | 0.11303 | 0.002784 | 0.00184 | 0.112354 | 0.105803 | 0.109708 | 0.112697 | 0.011843 | 0.013156 | 0.08928 | 0.096774 | 0.0926 | 0.110348 |  |  |  |  |  |  |  |
| 0.086334 | 0.088474 | 0.001605 | 0.001045 | 0.001281 | -0.00078 | 0.115677 | 0.111058 | 0.000677 | 0.003861 |  | 0.001042 | 0.094861 | 0.098574 | 0.003968 | 0.002576 | 0.004419 |  | 0.110428 |  |  |  |  |  |  |
| 0.086177 | 0.088964 | 0.004011 | 0.002181 | 0.00223 | 0.001023 | 0.113899 | 0.109214 | 0.000779 | 0.003207 |  | 0.001214 | 0.095642 | 0.098896 | 0.00524 | 0.003269 | 0.003833 |  | 0.10851 | 0.001189 |  |  |  |  |  |
| 0.092757 | 0.096123 | 0.003166 | 0.002844 |  | 0.002683 | 0.124765 | 0.123346 | 0.002544 | 0.012597 |  | 0.003245 | 0.103092 | 0.10509 | 0.006119 | 0.005004 | 0.00635 | 0.003414 | 0.119077 | 0.004873 | 0.004102 |  |  |  |  |
| 0.089095 | 0.091719 | 0.002625 |  |  |  | 0.119355 | 0.116868 |  | 0.004493 |  |  | 0.099969 | 0.103336 | 0.003105 | 0.001296 | 0.004239 |  | 0.114219 | 0.000939 | 0.001564 | 0.002349 |  |  |  |
| 0.068913 | 0.071459 | 0.006047 | 0.004493 | 0.004307 | 0.003797 | 0.099505 | 0.09529 | 0.004194 | 0.005332 | 0.004206 | 0.005062 | 0.079507 | 0.08279 |  |  | 4.92E-05 | 0.003701 | 0.094881 | 0.004693 | 0.005336 | 0.006856 | 0.003388 |  |  |
| 0.067422 | 0.069384 | 0.008331 | 0.007097 | 0.008009 | 0.00698 | 0.095056 | 0.090959 | 0.005991 | 0.006095 | 0.004841 | 0.006171 | 0.076465 | 0.080138 | 0.002415 | 0.002 | 0.00091 | 0.004639 | 0.091076 | 0.005089 | 0.006249 | 0.01347 |  | 0.000477 |  |
| 0.073109 | 0.076395 | 0.002059 | 0.001889 | 0.001155 |  | 0.102227 | 0.097743 | 0.001384 | 0.002467 | 0.000761 | 0.001721 | 0.08322 | 0.086502 |  |  |  |  | 0.097176 | 0.001475 | 0.001618 | 0.003281 |  |  | 0.001905 |

Table 1.3. A summary of relevant FVCOM models utilized in connectivity matrices.

| Year | Start Date (Julian) | End Date (Julian) | Vertical Velocity | Horizontal Diffusivity | \# Particles | Run Duration |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2014 | May 25 (145) | July 14 (195) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2014 | June 01 (152) | July 21 (202) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2014 | June 08 (159) | July 28 (209) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2014 | June 15 (166) | Aug 04 (216) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2014 | June 22 (173) | Aug 11 (223) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2014 | June 29 (180) | Aug 18 (230) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2015 | May 24 (144) | July 13 (194) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2015 | May 31 (151) | July 20 (201) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2015 | June 07 (158) | July 27 (208) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/ $/ \mathrm{s}$ | 224600 | 50 days |
| 2015 | June 14 (165) | Aug 03 (215) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2015 | June 21 (172) | Aug 10 (222) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2015 | June 28 (179) | Aug 17 (229) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2016 | May 29 (150) | July 18 (200) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2016 | June 05 (157) | July 25 (207) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/ | 224600 | 50 days |
| 2016 | June 12 (164) | Aug 01 (214) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2016 | June 19 (171) | Aug 08 (221) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/ $/ \mathrm{s}$ | 224600 | 50 days |
| 2016 | June 26 (178) | Aug 15 (228) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2016 | July 03 (185) | Aug 22 (235) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/ $/ \mathrm{s}$ | 224600 | 50 days |
| 2017 | May 28 (148) | July 17 (198) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2017 | June 04 (155) | July 24 (205) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m$/{ }^{2} / \mathrm{s}$ | 224600 | 50 days |
| 2017 | June 11 (162) | July 31 (212) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2017 | June 18 (169) | Aug 07 (219) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2017 | June 25 (176) | Aug 14 (226) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2017 | July 02 (183) | Aug 21 (233) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2018 | May 27 (147) | July 16 (197) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2018 | June 03 (154) | July 23 (204) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2018 | June 10 (161) | July 30 (211) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m$/{ }^{2} / \mathrm{s}$ | 224600 | 50 days |
| 2018 | June 17 (168) | Aug 06 (218) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m$/{ }^{2} / \mathrm{s}$ | 224600 | 50 days |
| 2018 | June 24 (175) | Aug 13 (225) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m$/{ }^{2} / \mathrm{s}$ | 224600 | 50 days |
| 2018 | July 01 (182) | Aug 20 (232) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2019 | May 26 (146) | July 15 (196) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/ | 224600 | 50 days |
| 2019 | June 02 (153) | July 22 (203) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2019 | June 09 (160) | July 29 (210) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/ $/ \mathrm{s}$ | 224600 | 50 days |
| 2019 | June 16 (167) | Aug 05 (217) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |
| 2019 | June 23 (174) | Aug 12 (224) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m/2/s | 224600 | 50 days |
| 2019 | June 30 (181) | Aug 19 (231) | $0.0003 \mathrm{~m} / \mathrm{s}$ | 5.6 m²/s | 224600 | 50 days |



Figure 1.2. (a) Isolation by distance (IBD) plot of $\mathrm{F}_{\text {ST }}$ values against geographic (Euclidean) distance among Green Bay populations. (b) Regression analysis of proportion Green Bay ancestry and distance from the mouth of Green Bay.

(a)
(b)


Figure 1.3. (a) Isolation by distance (IBD) plot of $\mathrm{F}_{\text {ST }}$ values against geographic (Euclidean) distance among main basin populations. (b) Regression analysis of genetic distance ( $\mathrm{F}_{\mathrm{ST}}$ ) and oceanic distance (1- dispersal probability), showing increased genetic distance with increased oceanic distance.

